

Time Evolution of SIS epidemics in the Complete Graph

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v1.: May 2013

v2: Oct 2016

v3: April 2017

Abstract

We show that, at the time being, the probability $\Pr[M(t) = k]$ that the number of infected nodes $M(t)$ at time t equals k in the Markovian continuous-time ε -SIS process on the complete graph cannot be determined exactly.

1 Introduction

In spite of the simplicity of the Markovian continuous-time SIS model, there does not seem to exist an exact time-dependent solution for any graph. Most analytic results are known for the complete graph as shown in [12, Sec. 17.6]. Before elaborating on the exact analytic solution of the Markovian continuous-time SIS model on the complete graph K_N containing N nodes, we briefly review the classical mean-field approximation.

For the complete graph K_N , mean-field approximations are accurate [3, 16]. Very likely – although there does not seem to be a rigorous proof – among all graphs, mean-field approximations are the most accurate in the complete graph. In the N -intertwined mean-field approximation (NIMFA) [15, 11], the governing equation for the probability $v(t)$ of infection in a node at time t in a regular graph G with degree r equals

$$\frac{dv(t)}{dt} = r\beta(t)v(t)(1-v(t)) - \delta(t)v(t) \quad (1)$$

where the infection rate $\beta(t)$ and the curing rate $\delta(t)$ are general non-negative real functions of time t . The probability $v(t)$ at time t changes due to two possible actions: (a) if the node is healthy with probability $1-v(t)$, its r infected neighbors – each neighbor is infected with the same probability $v(t)$ (due to symmetry) – can infect the node with instantaneous rate $\beta(t)$; (b) when the node is infected, which happens with probability $v(t)$, a curing processes with instantaneous rate $\delta(t)$ can heal the node. Since the rates are time-varying, the infection and curing process are independent, inhomogeneous Poisson processes [12]. The differential equation (1) can be solved exactly [13], resulting in

$$v(t) = \frac{\exp\left(\int_0^t (r\beta(u) - \delta(u)) du\right)}{\frac{1}{v_0} + r \int_0^t \beta(s) \exp\left(\int_0^s (r\beta(u) - \delta(u)) du\right) ds} \quad (2)$$

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where v_0 is the initial fraction of infected nodes.

As shown in [6] for regular graphs, the governing differential equations are precisely the same for NIMFA and the heterogeneous mean-field (HMF) approximation [8] of Pastor-Satorras and Vespignani. Hence, the equation (1) constitutes a general SIS mean-field approximation for regular graphs. An interesting feature of (1) is its independence on the size of the network, which avoids (or ignores) finite-size effects that often complicate studies of phase transitions. For regular graphs, the NIMFA average fraction of infected nodes $y(t) = v(t)$ and $y(t)$ is coined the order parameter in statistical physics. Equation (1) with constant rates, $\beta(t) = \beta$ and $\delta(t) = \delta$, has been investigated earlier by Kephart and White [5]. Many variations on and extensions of the epidemic Kephart and White model have been proposed (see e.g. [9, 15, 7]). In fact, the differential equation (1) with constant rates has already appeared in earlier work before Kephart and White (see e.g. [2, 4]) and is also known as the logistic differential equation of population growth, first introduced by Verhulst [17] in 1845.

2 The number of infected nodes in K_N

We consider the time-dependent ε -SIS process on the complete graph, where a positive self-infection rate ε is crucial for the existence of a non-trivial steady state as shown in [12, Chapter 17]. The number of infected nodes $M(t)$ at time t in the complete graph K_N is described by a continuous-time Markov process on $\{0, 1, \dots, N\}$ with the following rates:

$$\begin{aligned} M &\mapsto M + 1 \text{ at rate } (\beta M + \varepsilon)(N - M) \\ M &\mapsto M - 1 \text{ at rate } \delta M. \end{aligned}$$

Every infected node heals with rate δ , which explains the transition rate $M \mapsto M - 1$. Every healthy node (of which there are $N - M$ at state M) has exactly M infected neighbors, each actively transferring the virus with rate β in addition to the self-infection rate ε . Alternatively, each of the M infected nodes can infect its $N - M$ healthy neighbors with a rate β and the $N - M$ healthy nodes can infect themselves with self-infection rate ε .

This Markov process $M(t)$ is a birth and death process with birth rate $\lambda_k = (\beta k + \varepsilon)(N - k)$ and death rate $\mu_k = k\delta$ when it is in a state with $M(t) = k$ infected nodes. The steady-state probabilities π_0, \dots, π_N , where $\pi_k = \lim_{t \rightarrow \infty} \Pr[M(t) = k]$, of a general birth-death process can be computed exactly [12, p. 230],[14] as

$$\pi_k = \pi_0 \binom{N}{k} \varepsilon^* \tau^{k-1} \frac{\Gamma(\frac{\varepsilon^*}{\tau} + k)}{\Gamma(\frac{\varepsilon^*}{\tau} + 1)} \quad (k > 0) \quad (3)$$

and

$$\pi_0 = \frac{1}{\sum_{k=0}^N \binom{N}{k} \tau^k \frac{\Gamma(\frac{\varepsilon^*}{\tau} + k)}{\Gamma(\frac{\varepsilon^*}{\tau})}} \quad (4)$$

where the effective infection rate $\tau = \frac{\beta}{\delta}$ and $\varepsilon^* = \frac{\varepsilon}{\delta}$. Thus, π_0 is the steady-state probability that the complete graph K_N is infection free or overall healthy. When $\varepsilon \rightarrow 0$ for N fixed, we observe from (3) that $\lim_{\varepsilon \rightarrow 0} \pi_k = 0$ for $k > 0$ and, consequently, that $\lim_{\varepsilon \rightarrow 0} \pi_0 = 1$, which reflects that the steady state of the SIS process (in any finite graph) is the overall-healthy state or absorbing state.

3 A generating function approach

We denote the probability that the number of infected nodes $M(t)$ at time t equals k (or that the ε -SIS process at time t is in state k) by

$$s_k(t) = \Pr[M(t) = k] \quad (5)$$

By convention, we agree that $s_k(t) = 0$ if $k > N$ or if $k < 0$. Thus, $s_0(t)$ is the probability that the epidemic dies out at time t or that the complete graph K_N is infection free at time t , but only remains infection free provided the self-infection rate $\varepsilon = 0$. Further, the steady-state probabilities

$$\pi_k = \lim_{t \rightarrow \infty} s_k(t)$$

are explicitly known in (3). The birth rate $\lambda_k = (\beta k + \varepsilon)(N - k) = -\beta k^2 + (N\beta - \varepsilon)k + N\varepsilon$ is quadratic in k and the death rate $\mu_k = \delta k$ is linear in k for any state $k \in \{0, 1, 2, \dots, N\}$. The time-dependent evolution of the constant birth and death process [12, p. 239] as well as the linear birth and death process is described in [12, p. 243]. Here, we study the quadratic birth and death process, whose solution has, by the best of our efforts, not yet appeared in the literature.

Applying the differential equations of a general birth and death process to ε -SIS process yields the set

$$s'_0(t) = \delta s_1(t) - N\varepsilon s_0(t) \quad (6)$$

$$s'_k(t) = \{\beta k^2 - (N\beta + \delta - \varepsilon)k - N\varepsilon\} s_k(t) + \{-\beta(k-1)^2 + (N\beta - \varepsilon)(k-1) + N\varepsilon\} s_{k-1}(t) + \delta(k+1) s_{k+1}(t) \quad (7)$$

where all involved rates β, δ and ε can depend upon time t . The first differential equation (6) is incorporated in the general one (7) for $k = 0$, since $s_{-1}(t) = 0$ by our convention. If $k = N$, then $\lambda_N = 0$ as well as $s_{N+1}(t)$, so that (7) reduces to

$$s'_N(t) = -\delta N s_N(t) + \{\beta(N-1) + \varepsilon\} s_{N-1}(t)$$

Since the ε -SIS epidemic must always be in one of the possible states, there holds that $\sum_{k=0}^N s_k(t) = 1$.

Following the general method illustrated in [12, Sec. 11.3.3-11.3.4] for the constant and linear rate birth and death process, we start by defining the probability generating function (pgf)

$$\varphi(x, t) = E[x^{M(t)}] = \sum_{k=0}^N s_k(t) x^k \quad (8)$$

which we can equally well write as $\varphi(x, t) = \sum_{k=0}^{\infty} s_k(t) x^k$, according to the convention that $s_k(t) = 0$ if $k > N$ or if $k < 0$. For any probability generating function $\varphi_X(z) = E[z^X] = \sum_{k=0}^{\infty} \Pr[X = k] z^k$, the radius R of convergence around $z = 0$ in the complex z -plane is at least equal to one, because for $|z| \leq 1$, it holds that $|\varphi_X(z)| \leq \sum_{k=0}^{\infty} \Pr[X = k] |z|^k \leq \sum_{k=0}^{\infty} \Pr[X = k] = \varphi_X(1) = 1$.

Theorem 1 *In the time-dependent ε -SIS process on the complete graph K_N , the probability generating function $\varphi(x, t)$ of the number of infected nodes $M(t)$ at time t obeys the partial differential equation*

$$\frac{\partial \varphi}{\partial t} = (x-1) \left\{ -\beta x^2 \frac{\partial^2 \varphi}{\partial x^2} + \{[(N-1)\beta - \varepsilon]x - \delta\} \frac{\partial \varphi}{\partial x} + N\varepsilon \varphi \right\} \quad (9)$$

Proof: After multiplying both sides in (7) by x^k and summing over all $k \geq 0$, the first line in (7) is transformed as

$$T_1 = \sum_{k=0}^N \{ \beta k^2 - (N\beta + \delta - \varepsilon)k - N\varepsilon \} s_k(t) x^k$$

With $\frac{\partial \varphi}{\partial x} = \sum_{k=0}^N k s_k(t) x^{k-1}$ and $\frac{\partial^2 \varphi}{\partial x^2} = \sum_{k=0}^N k(k-1) s_k(t) x^{k-2}$, we have

$$T_1 = \beta x^2 \frac{\partial^2 \varphi}{\partial x^2} - ((N-1)\beta + \delta - \varepsilon) x \frac{\partial \varphi}{\partial x} - N\varepsilon \varphi$$

Similarly, the transform of the second line in (7) taking our convention $s_{-1}(t) = 0$ into account is

$$T_2 = \sum_{k=1}^N \left\{ -\beta(k-1)^2 + (N\beta - \varepsilon)(k-1) + N\varepsilon \right\} s_{k-1}(t) x^k$$

leading to

$$T_2 = -\beta x^3 \frac{\partial^2 \varphi}{\partial x^2} + ((N-1)\beta - \varepsilon) x^2 \frac{\partial \varphi}{\partial x} + N\varepsilon x \varphi$$

Finally, the transform of the third and last line in (7) is, with $s_{N+1}(t) = 0$,

$$T_3 = \delta \sum_{k=0}^N (k+1) s_{k+1}(t) x^k = \delta \sum_{k=1}^{N+1} k s_k(t) x^{k-1} = \delta \sum_{k=1}^N k s_k(t) x^{k-1} = \delta \frac{\partial \varphi}{\partial x}$$

Equating the three right-hand side contributions $T_1 + T_2 + T_3$ and the transform of the left-hand side in (7) yields

$$\frac{\partial \varphi}{\partial t} = -\beta x^2 (x-1) \frac{\partial^2 \varphi}{\partial x^2} + \{ (N-1)\beta x(x-1) - \varepsilon x(x-1) - \delta(x-1) \} \frac{\partial \varphi}{\partial x} + N\varepsilon (x-1) \varphi$$

Thus, we find the partial differential equation (9). \square

The factor $(x-1)$ at the right-hand side of (9) is a consequence of the conservation of probability at any time t , namely that $\varphi(1, t) = \sum_{k=0}^{\infty} s_k(t) = 1$, implying that the ε -SIS stochastic process is surely in one of the possible states. Furthermore, $\left. \frac{\partial \varphi}{\partial x} \right|_{x=1} = \sum_{k=0}^{\infty} k s_k(t)$ is the average number of infected nodes at time t . Hence, the average fraction of infected nodes at time t equals

$$y(t; \tau) = \frac{1}{N} \left. \frac{\partial \varphi(x, t)}{\partial x} \right|_{x=1} \quad (10)$$

Initial condition. The ε -SIS process can start with a certain probability distribution, which then requires that the initial state vector $s(0) = (s_0(0), s_1(0), \dots, s_N(0))$ is given. When precisely m nodes in K_N are infected initially at $t = 0$, then the boundary condition $\varphi(x, 0) = \sum_{k=0}^{\infty} \delta_{km} x^k = x^m$. Clearly, the value of $m > 0$ must exceed zero, because $\varphi(0, t) = s_0(t)$ is the probability that the complete graph is infection free at time t and, on the long run, $\lim_{t \rightarrow \infty} \varphi(0, t) = \pi_0$ is given by (4).

Confinement. *In the sequel, we limit ourselves to constant rates: none of the infection rate β , self-infection rate ε or curing rate δ is a function of time t . In addition, we assume that the ε -SIS process starts at $t = 0$.*

3.1 The steady-state probability generating function $\varphi_\infty(x)$

The steady-state probability generating function (assuming constant rates) equals with (3)

$$\lim_{t \rightarrow \infty} \varphi(x, t) = \sum_{k=0}^{\infty} \lim_{t \rightarrow \infty} \Pr [M(t) = k] x^k = \sum_{k=0}^{\infty} \pi_k x^k = \varphi_\infty(x)$$

where

$$\varphi_\infty(x) = \pi_0 + \frac{\varepsilon^* \pi_0}{\tau \Gamma\left(\frac{\varepsilon^*}{\tau} + 1\right)} \sum_{k=1}^N \binom{N}{k} \Gamma\left(\frac{\varepsilon^*}{\tau} + k\right) (\tau x)^k \quad (11)$$

Thus, if $\varepsilon = 0$, then $\pi_0 = 1$ and there holds that $\lim_{t \rightarrow \infty} \varphi(x, t) = \varphi_\infty(x) = 1$. If $\varepsilon > 0$, the steady-state probability generating function $\varphi_\infty(x)$ is a polynomial of degree N in x , which is more elegantly written as

$$\varphi_\infty(x) = \frac{\pi_0}{\Gamma\left(\frac{\varepsilon^*}{\tau}\right)} \sum_{k=0}^N \binom{N}{k} \Gamma\left(\frac{\varepsilon^*}{\tau} + k\right) (\tau x)^k \quad (12)$$

and the general relation for any pgf, $\varphi_\infty(1) = 1$, also follows from (4). Finally, $\varphi_\infty(x)$ is a function of three parameters

$$\varphi_\infty(x) = \varphi_\infty(x; \tau, \varepsilon^*, N)$$

The partial differential equation (9) simplifies, in the steady state for $t \rightarrow \infty$ and $\frac{\partial \varphi}{\partial t} = 0$, to

$$-\beta x^2 \frac{\partial^2 \varphi_\infty}{\partial x^2} + \{[(N-1)\beta - \varepsilon]x - \delta\} \frac{\partial \varphi_\infty}{\partial x} + N\varepsilon \varphi_\infty = 0 \quad (13)$$

Introducing the integral for the Gamma function $\Gamma(s) = \int_0^\infty u^{s-1} e^{-u} du$, valid for $\text{Re}(s) > 0$, into (12) yields

$$\begin{aligned} \varphi_\infty(x) &= \frac{\pi_0}{\Gamma\left(\frac{\varepsilon^*}{\tau}\right)} \sum_{k=0}^N \binom{N}{k} (\tau x)^k \int_0^\infty u^{\frac{\varepsilon^*}{\tau} + k - 1} e^{-u} du \\ &= \frac{\pi_0}{\Gamma\left(\frac{\varepsilon^*}{\tau}\right)} \int_0^\infty u^{\frac{\varepsilon^*}{\tau} - 1} e^{-u} \left\{ \sum_{k=0}^N \binom{N}{k} (u\tau x)^k \right\} du \end{aligned}$$

Invoking Newton's binomium leads to an integral representation¹ of the steady-state probability generating function for $\varepsilon > 0$,

$$\varphi_\infty(x; \tau, \varepsilon^*, N) = \frac{\pi_0}{\Gamma\left(\frac{\varepsilon^*}{\tau}\right)} \int_0^\infty u^{\frac{\varepsilon^*}{\tau} - 1} e^{-u} (1 + u\tau x)^N du \quad (15)$$

¹ Assuming a positive real x and letting $w = (\tau x)u$, we find

$$\varphi_\infty(x) = \frac{\pi_0}{(\tau x)^{\frac{\varepsilon^*}{\tau}} \Gamma\left(\frac{\varepsilon^*}{\tau}\right)} \int_0^\infty e^{-\frac{w}{\tau x}} w^{\frac{\varepsilon^*}{\tau} - 1} (1 + w)^N dw$$

We conclude that the steady-state probability generating function $\varphi_\infty(x)$ can be written as

$$\varphi_\infty(x) = \frac{\pi_0}{(\tau x)^{\frac{\varepsilon^*}{\tau}}} U\left(\frac{\varepsilon^*}{\tau}, \frac{\varepsilon^*}{\tau} + 1 + N, \frac{1}{\tau x}\right) \quad (14)$$

where the confluent hypergeometric function [1, 13.2.8]

$$U(a, b, z) = \frac{1}{\Gamma(a)} \int_0^\infty e^{-zt} w^{a-1} (1+w)^{b-a-1} dt$$

is one of the independent solutions of Kummer's differential equation $x \frac{d^2 f}{dx^2} + (b-x) \frac{df}{dx} - af = 0$ (see e.g. [1, Chapter 13]).

3.2 General solution of the partial differential equation (9)

Theorem 2 *In the time-dependent ε -SIS process on the complete graph K_N with constant infection rate β , self-infection rate ε and curing rate δ , the probability generating function $\varphi(x, t)$ of the number of infected nodes $M(t)$ at time t can be written as a Laplace transform*

$$\varphi(x, t) = \int_0^\infty e^{-ct} g(x; c) dc \quad (16)$$

where the function $g(x, c)$ obeys the differential equation

$$-x^2(x-1) \frac{d^2g}{dx^2} + \left\{ \left[(N-1) - \frac{\varepsilon^*}{\tau} \right] x - \frac{1}{\tau} \right\} (x-1) \frac{dg}{dx} + \frac{1}{\tau} (N\varepsilon^*(x-1) + c^*) g = 0 \quad (17)$$

and $\tau = \frac{\beta}{\delta}$, $\varepsilon^* = \frac{\varepsilon}{\delta}$ and $c^* = \frac{c}{\delta} \geq 0$.

Proof: The usual recipe of the separation of the variables t and x , by assuming that a solution in product form as $\varphi(x, t) = g(x) h(t)$ exists, transforms (9) to

$$\frac{\partial \log h}{\partial t} = \frac{(x-1)}{g} \left\{ -\beta x^2 \frac{d^2g}{dx^2} + \{[(N-1)\beta - \varepsilon]x - \delta\} \frac{dg}{dx} + N\varepsilon g \right\} \quad (18)$$

By taking the derivative of both sides with respect to x , we find with $\frac{\partial}{\partial x} \frac{\partial \log h}{\partial t} = 0$ that

$$\frac{(x-1)}{g} \left\{ -\beta x^2 \frac{d^2g}{dx^2} + \{[(N-1)\beta - \varepsilon]x - \delta\} \frac{dg}{dx} + N\varepsilon g \right\} = c_1 \quad (19)$$

where c_1 is a constant that is neither a function of x nor of t , because the left-hand side in (19) is independent of t . Similarly, by taking the derivative of both sides in (18) with respect to t , we find that

$$\frac{\partial \log h}{\partial t} = c_2 \quad (20)$$

and (18) shows that $c_1 = c_2 = -c$.

We rewrite (19) with $\varepsilon^* = \frac{\varepsilon}{\delta}$ and $c^* = \frac{c}{\delta}$ to find (17).

From (20), we find $h(t) = h(0) e^{-ct}$ for the time $t \geq 0$. If c were complex and $\text{Im}(c) \neq 0$, then $h(t) = h(0) e^{-\text{Re}(c)t} (\cos t \text{Im}(c) + i \sin t \text{Im}(c))$ and $\varphi(x, t) = g(x) h(t)$ is generally complex for $t > 0$. However, the definition (8) of the pgf $\varphi(x, t)$ illustrates that $\varphi(x, t)$ is real for real x at any time $t \geq 0$. Hence, c must be real. Moreover, since the asymptotic pgf $\lim_{t \rightarrow \infty} \varphi(x, t) = \varphi_\infty(x)$ exists, c must be non-negative, otherwise $\lim_{t \rightarrow \infty} h(t) = h(0) \lim_{t \rightarrow \infty} e^{-ct} = \infty$. We conclude that the eigenvalue c is real and non-negative.

The general solution of the eigenvalue differential equation in c consists of a linear combination $\sum_{c \geq 0} e^{-ct} g(x; c)$ if the eigenvalues c are discrete. Generally, one readily verifies that $\varphi(x, t) = \int_0^\infty e^{-ct} g(x; c) dc$ satisfies the partial differential equation (9) provided that $g(x; c)$ is a solution of the differential equation (17) as a function of the ‘‘eigenvalue’’ c . \square

In fact, we need to solve an eigenvalue problem that can be expanded in a Sturm-Liouville series [10]. For $c = 0$, the differential equation (17) reduces to the differential (13) and we conclude that

$$g(x, 0) = \varphi_\infty(x)$$

The ε -SIS process on the complete graph K_N with N nodes is described by a general birth-death process by the differential equations (6) and (7). This set of linear differential equations possesses a general $(N + 1) \times (N + 1)$ tri-diagonal matrix, whose eigenstructure is studied in depth in [12, A.6.3]. The $N + 1$ non-negative, real eigenvalues (and one of them is zero) imply that the eigenvalues c are a discrete set $\{c_0 = 0, c_1, \dots, c_N\}$, so that the Laplace integral in (16) will reduce to a sum $\varphi(x, t) = \varphi_\infty(x) + \sum_{k=1}^N e^{-c_k t} g(x; c_k)$ for finite size N .

The second-order differential equation (17) in the function g is of the type

$$x^2 (1 - x) \frac{d^2 g}{dx^2} + (ax + b) (1 - x) \frac{dg}{dx} + (\lambda + d(1 - x)) g = 0 \quad (21)$$

where $a = \frac{\varepsilon^*}{\tau} - (N - 1)$, $b = \frac{1}{\tau}$, $d = N \frac{\varepsilon^*}{\tau}$ and $\lambda = \frac{c^*}{\tau}$ are real numbers. Unfortunately, (21) does not seem to be of a known type. Gauss's hypergeometric function $F(a, b; c; x)$ obeys the differential equation [1, Chapter 15]

$$x(1-x) \frac{d^2 g}{dx^2} + [c - (a+b+1)x] \frac{dg}{dx} - abg = 0$$

Slightly more general, (17) is of the type

$$p_3(x) g^{(2)}(x) + p_2(x) g^{(1)}(x) + p_1(x) g(x) = 0$$

where $p_k(x)$ is a polynomial in x of degree k , where the hypergeometric differential equation is of the form

$$p_2(x) g^{(2)}(x) + p_1(x) g^{(1)}(x) + p_0(x) g(x) = 0$$

In conclusion, unless an analytic solution of the differential (21) can be found, we are afraid that the probability $s_k(t) = \Pr[M(t) = k]$, that the number of infected nodes $M(t)$ at time t equals k in the Markovian continuous-time ε -SIS process on the complete graph K_N , cannot be determined exactly.

Acknowledgement. I am grateful to Johan Dubbeldam for checking the computations and for useful discussions.

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A Reduction of the differential equation (17) to the standard form

We aim to transform (21) into the form [10]

$$\frac{d^2 y}{du^2} + (\lambda - q(u)) y(u) = 0 \quad (22)$$

The standard form has many interesting properties. First, the Wronskian is constant in u . Second, Titchmarsh [10] gives, at the beginning of the chapters, insight in the spectrum of λ and he also presents bounds to the solution y .

We make the transformation $x = h(u)$, so that $u = h^{-1}(x)$. Thus, by using the chain rule and denoting $f(u) = g(h(u))$, we have

$$\frac{dg(x)}{dx} = \frac{d}{du} g(h(u)) \frac{du}{dx} = \frac{df}{du} \frac{1}{\frac{dx}{du}} = \frac{1}{h'(u)} \frac{df}{du}$$

and

$$\frac{d^2 g(x)}{dx^2} = \frac{d}{dx} \left(\frac{dg(x)}{dx} \right) = \frac{d}{du} \left(\frac{dg(x)}{dx} \right) \frac{du}{dx} = \frac{1}{(h'(u))^2} \frac{d^2 f}{du^2} - \frac{h''(u)}{(h'(u))^3} \frac{df}{du}$$

We obtain

$$\frac{h^2(1-h)}{(h'(u))^2} \frac{d^2 f}{du^2} + \frac{(1-h)}{h'(u)} \left\{ (ah+b) - h^2 \frac{h''(u)}{(h'(u))^2} \right\} \frac{df}{du} + (\lambda + d(1-h)) f = 0 \quad (23)$$

Nex, we choose h such that $h^2(1-h) \frac{1}{(h'(u))^2} = 1$. Thus, $h^2(1-h) = (h'(u))^2$ or $\frac{dh}{du} = \pm h\sqrt{1-h}$ and integrated

$$\pm \int \frac{dh}{h\sqrt{1-h}} = u$$

As in Tichmarsh [10], we assume the positive sign and find

$$u = \log \frac{1 - \sqrt{1-h}}{1 + \sqrt{1-h}}$$

and, inversed,

$$x = h(u) = \operatorname{sech}^2\left(\frac{u}{2}\right)$$

Thus, $x = h(u) = \operatorname{sech}^2\left(\frac{u}{2}\right)$ and $u = 2\operatorname{ArcSech}(\sqrt{x})$, which is only real and positive for $x \in (0, 1)$.

After introducing $h(u) = \operatorname{sech}^2\left(\frac{u}{2}\right)$ into (23) yields

$$\frac{d^2 f}{du^2} + \left\{ \frac{\cosh u - 2}{\sinh u} - a \tanh \frac{u}{2} - \frac{b}{2} \sinh u \right\} \frac{df}{du} + \left(\lambda + d \tanh^2 \frac{u}{2} \right) f = 0$$

Let

$$r(u) = \frac{\cosh u - 2}{\sinh u} - a \tanh \frac{u}{2} - \frac{b}{2} \sinh u \quad (24)$$

then we obtain the differential equation in $f(u) = g(\operatorname{sech}^2\left(\frac{u}{2}\right))$ and $x = \operatorname{sech}^2\left(\frac{u}{2}\right)$ or $u = 2\operatorname{ArcSech}(\sqrt{x})$,

$$\frac{d^2 f}{du^2} + r(u) \frac{df}{du} + \left(\lambda + d \tanh^2 \frac{u}{2} \right) f = 0$$

We proceed with the reduction to the standard form by considering $f(u) = p(u)s(u)$ and the above differential equation becomes

$$0 = p''(u) + p'(u) \left\{ 2 \frac{s'(u)}{s(u)} + r(u) \right\} + p(u) \left\{ \frac{s''(u)}{s(u)} + r(u) \frac{s'(u)}{s(u)} + \left(\lambda + d \tanh^2 \frac{u}{2} \right) \right\}$$

The standard form requires that $2\frac{s'(u)}{s(u)} + r(u) = 0$, or

$$2\frac{d}{du} \log s(u) = -r(u)$$

and

$$s(u) = \exp\left(-\frac{1}{2} \int r(u) du\right)$$

Explicitly, we have

$$s(u) = \exp\left(-\frac{1}{2} \int r(u) du\right) = \frac{\tanh u \left(\cosh \frac{u}{2}\right)^a}{\sqrt{\sinh u}} e^{\frac{b}{4} \cosh u} \quad (25)$$

From $2\frac{s'(u)}{s(u)} + r(u) = 0$ or $\frac{s'(u)}{s(u)} = -\frac{1}{2}r(u)$, equivalent to $2s'(u) + s(u)r(u) = 0$, we find that

$$s''(u) = -\frac{1}{2}s'(u)r(u) - \frac{1}{2}s(u)r'(u)$$

which we use in

$$X = \frac{s''(u)}{s(u)} + r(u)\frac{s'(u)}{s(u)} = -\frac{1}{4}\{r^2(u) + 2r'(u)\}$$

Hence, with $s(u)$ in (25) and obeying $\frac{s'(u)}{s(u)} = -\frac{1}{2}r(u)$ and with $p(u) = \frac{f(u)}{s(u)}$, we arrive at

$$p''(u) + p(u) \left\{ \lambda + d \tanh^2 \frac{u}{2} - \frac{1}{4} \{r^2(u) + 2r'(u)\} \right\} = 0$$

so that

$$q(u) = \frac{1}{4} \{r^2(u) + 2r'(u)\} - d \tanh^2 \frac{u}{2} \quad (26)$$

We now compute $q(u)$. From the definition (24) of $r(u)$,

$$\begin{aligned} r^2(u) + 2r'(u) &= 1 - a + a(a+1) \tanh^2 \frac{u}{2} + b \left\{ \frac{b}{4} \cosh u + \frac{b}{4} - 2 \right\} (\cosh u - 1) + \frac{1}{\sinh^2 u} \\ &\quad + a \tanh \frac{u}{2} \left(b \sinh u - 2 \frac{\cosh u - 2}{\sinh u} \right) \end{aligned}$$

which is not such an insightfull expression!

Finally, we arrive with $p(u) = \frac{f(u)}{s(u)}$ at the standard form

$$p''(u) + \left(\lambda + d \tanh^2 \frac{u}{2} - \frac{1}{4} \{r^2(u) + 2r'(u)\} \right) p(u) = 0$$

Explicitly, with $\lambda' = \lambda + \frac{a}{4} - \frac{b}{2}$

$$p''(u) + \left(\lambda' + \left(d - \frac{a^2}{4} - \frac{a}{4} \right) \tanh^2 \frac{u}{2} - \frac{b^2}{16} \sinh^2 u - \frac{\cosh^2 u + 2}{4 \sinh^2 u} - \frac{ab}{4} \sinh u \tanh \frac{u}{2} \right. \\ \left. + \frac{1}{2} a \tanh \frac{u}{2} \frac{\cosh u - 2}{\sinh u} + \frac{b}{2} \cosh u \right) p(u) = 0$$